

# Novel Spectrum-Sensing Algorithms for Cognitive Radio Based on with and without Signal Information

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## To Cite this Article

K.Nayomi and M.Sailaja, "Novel Spectrum-Sensing Algorithms for Cognitive Radio Based on with and without Signal Information", *International Journal for Modern Trends in Science and Technology*, Vol. 03, Issue 01, 2017, pp. 60-66.

## ABSTRACT

*In first module presents a spectrum monitoring algorithm for Orthogonal Frequency Division Multiplexing (OFDM) based cognitive radios by which the primary user reappearance can be detected during the secondary user transmission. The proposed technique reduces the frequency with which spectrum sensing must be performed and greatly decreases the elapsed time between the start of a primary transmission and its detection by the secondary network. In second module the statistical covariance of the received signal and noise are usually different, they can be used to differentiate the case where the primary user's signal is present from the case where there is only noise. In second module, spectrum-sensing algorithms are proposed based on the sample covariance matrix calculated from a limited number of received signal samples. Two test statistics are then extracted from the sample covariance matrix. A decision on the signal presence is made by comparing the two test statistics. Theoretical analysis for the proposed algorithms is given. Detection probability and the associated threshold are found based on the statistical theory. The methods do not need any information about the signal, channel, and noise power a priori. In addition, no synchronization is needed. Simulations based on narrow-band signals, captured digital television (DTV) signals, and multiple antenna signals are presented to verify the methods.*

**KEYWORDS:** Communication, communication channels, covariance matrices, signal detection, signal processing.

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## I. INTRODUCTION

Nowadays, static spectrum access is the main policy for wireless communications. Under this policy, fixed channels are assigned to licensed users or primary users (PUs) for exclusive use while unlicensed users or secondary users (SUs) are prohibited from accessing those channels even when they are unoccupied. The idea of a cognitive radio (CR) was proposed to achieve more efficient utilization of the RF spectrum [1]. One of the main approaches utilized by cognitive networks is the

overlay network model [2] in which SUs seek to opportunistically use the spectrum when the PUs are idle. Primary and secondary users are not allowed to operate simultaneously. In this method, secondary users must sense the spectrum to detect whether it is available or not prior to communication. If the PU is idle, the SU can then use the spectrum, but it must be able to detect very weak signals from the primary user by monitoring the shared band to quickly vacate the occupied spectrum. During this process, the CR system may spend a long time, known as the

sensing interval, during which the secondary transmitters are silent while the frequency band is sensed. Since the CR users do not utilize the spectrum during the detection time, these periods are also called quiet periods (QPs) [3]. In the IEEE 802.22 system, a quiet period consists of a series of consecutive spectrum sensing intervals using energy detection to determine if the signal level is higher than a predefined value, which indicates a non-zero probability of primary user transmission. The energy detection is followed by feature detection to distinguish whether the source of energy is a primary user or noise [4], [5]. This mechanism is repeated periodically to monitor the spectrum. Once the PU is detected, the SU abandons the spectrum for a finite period and chooses another valid spectrum band in the spectrum pool for communication.

OFDM is a multi-carrier modulation technique that is used in many wireless systems and proven as a reliable and effective transmission method. For these reasons, OFDM is utilized as the physical layer modulation technique for many wireless systems including DVB-T/T2, LTE, IEEE 802.16d/e, and IEEE 802.11a/g. Similar to other wireless networks, OFDM is preferred for cognitive networks and has been already in use for the current cognitive standard IEEE 802.22. On the other hand, OFDM systems have their own challenges that need special treatment [10]. These challenges include its sensitivity to frequency errors and the large dynamic range of the time domain signal. Moreover, the finite time-window in the receiver DFT results in a spectral leakage from any in-band and narrow band signal onto all OFDM sub-carriers.

The paper is organized as follows. Section II summarizes the overall system model. In Section III, the *energy ratio* technique is discussed. In addition, we present performance analysis for AWGN channels under perfect synchronization and neglecting power leakage in Section IV. OFDM challenges and non-perfect synchronization environment are considered in Section V. Further, we extend the analysis and study to frequency selective fading channels and multi-antenna systems in Section VI. The complexity of the *energy ratio* is analyzed and architecture is also proposed in Section VII. Finally, the performance is evaluated with computer simulations in Section VIII.

## II. COVARIANCE-BASED DETECTIONS

Let  $x_c(t) = s_c(t) + \eta_c(t)$  be the continuous-time received signal, where  $s_c(t)$  is the possible primary users signal and  $\eta_c(t)$  is the noise.  $\eta_c(t)$  is assumed to be a stationary process satisfying  $E(\eta_c(t)) = 0$ , and  $E(\eta_c^2(t)) = \sigma_\eta^2$ , and  $E(\eta_c(t)\eta_c(t+\tau)) = 0$  for any  $\tau \neq 0$ . Assume that we are interested in the frequency band with central frequency  $f_c$  and bandwidth  $W$ . We sample the received signal at a sampling rate  $f_s$  where  $f_s \geq W$

Let  $T_s = \frac{1}{f_s}$  be the sampling period. For notations simplicity, we define  $x(n) \square x_c(nT_s)$ ,  $s(n) \square s_c(nT_s)$ , and  $\eta(n) \square \eta_c(nT_s)$ . There are two hypotheses: 1)  $H_0$ , i.e., the signal does not exist, and 2)  $H_1$ , i.e., the signal exists. The received signal samples under the two hypotheses are given by [11], [12], [15], and [16]

$$H_0 : x(n) = \eta(n) \quad (1)$$

$$H_1 : x(n) = s(n) + \eta(n) \quad (2)$$

respectively, where  $s(n)$  is the transmitted signal samples that passed through a wireless channel consisting of path loss, multipath fading, and time dispersion effects; and  $\eta(n)$  is the white noise, which is i.i.d., having mean zero and variance  $\sigma_\eta^2$

Note that  $s(n)$  can be the superposition of the received signals from multiple primary users. No synchronization is needed here.

Two probabilities are of interest for spectrum sensing: 1) probability of detection  $P_d$  which defines, at hypothesis  $H_1$ , the probability of the sensing algorithm having detected the presence of the primary signal, and 2) probability of false alarm  $P_{fa}$  which defines, at hypothesis  $H_0$ , the probability of the sensing algorithm claiming the presence of the primary signal.

A. CAV Detection

Let us consider L consecutive samples and define the following vectors:

$$x(n) = [x(n) \quad x(n-1) \quad \dots \quad x(n-L+1)]^T$$

$$s(n) = [s(n) \quad s(n-1) \quad \dots \quad s(n-L+1)]^T$$

$$\eta(n) = [\eta(n) \quad \eta(n-1) \quad \dots \quad \eta(n-L+1)]^T$$

Parameter L is called the smoothing factor in the following. Considering the statistical covariance matrices of the signal and noise defined as

$$R_x = E[x(n)x^T(n)] \tag{6}$$

$$R_s = E[s(n)s^T(n)] \tag{7}$$

We can verify that

$$R_x = R_s + \sigma_\eta^2 I_L \tag{8}$$

If signal  $s(n)$  is not present,  $R_s = 0$ . Hence, the off-diagonal elements of  $R_x$  are all zeros. If there is a signal and the signal samples are correlated,  $R_s$  is not a diagonal matrix. Hence, some of the off-diagonal elements of  $R_x$  should be nonzeros.

Denote  $r_{nm}$  as the element of matrix  $R_x$  at the nth row and mth column, and let

$$T_1 = \frac{1}{L} \sum_{n=1}^L \sum_{m=1}^L |r_{nm}| \tag{9}$$

$$T_2 = \frac{1}{L} \sum_{n=1}^L |r_{nn}| \tag{10}$$

Then, if there is no signal,  $T_1/T_2 = 1$ . If the signal is present,  $T_1/T_2 > 1$ . Hence, ratio  $T_1/T_2$  can be used to detect the presence of the signal.

In practice, the statistical covariance matrix can only be calculated using a limited number of signal samples. Define the sample autocorrelations of the received signal as

$$\lambda(l) = \frac{1}{N_s} \sum_{m=0}^{N_s-1} x(m)x(m-l), \quad l = 0, 1, \dots, L-1$$

where  $N_s$  is the number of available samples. Statistical covariance matrix  $R_x$  can be approximated by the sample covariance matrix defined as

$$\hat{R}_x(N_s) = \begin{bmatrix} \lambda(0) & \lambda(1) & \dots & \lambda(L-1) \\ \lambda(1) & \lambda(0) & \dots & \lambda(L-2) \\ \vdots & \vdots & \ddots & \vdots \\ \lambda(L-1) & \lambda(L-2) & \dots & \lambda(0) \end{bmatrix}$$

Note that the sample covariance matrix is symmetric and Toeplitz. Based on the sample covariance matrix, we propose the following signal detection method:

Algorithm 1: Covariance Absolute Value (CAV) Detection Algorithm

Step 1) Sample the received signal, as previously described.

Step 2) Choose a smoothing factor  $L$  and a threshold  $\gamma_1$ , where  $\gamma_1$  should be

Chosen to meet the requirement for the probability of false alarm. This will be discussed in the next section.

Step 3) Compute the autocorrelations of the received signal  $\lambda(l), l = 0, 1, \dots, L-1$ , and form the sample co-variance matrix.

Step 4) compute

$$T_1(N_s) = \frac{1}{L} \sum_{n=1}^L \sum_{m=1}^L |r_{nm}(N_s)| \tag{13}$$

$$T_2(N_s) = \frac{1}{L} \sum_{n=1}^L |r_{nn}(N_s)| \tag{14}$$

Where  $r_{nm}(N_s)$  are the elements of the sample covariance matrix  $\hat{R}_x(N_s)$ .

Step 5) Determine the presence of the signal based on  $T_1(N_s)$ ,  $T_2(N_s)$ , and threshold  $\gamma_1$ . That is, if

$T_1(N_s)/T_2(N_s) > \gamma_1$ , the signal exists; otherwise, the signal does not exist.

B. Theoretical Analysis for the CAV Algorithm

The proposed method only uses the received signal samples. It does not need any information of the signal, channel, and noise power *a priori*. In addition, no synchronization is needed. The validity of the proposed CAV algorithm relies on the assumption that the signal samples are correlated, i.e.,  $R_s$  is not a diagonal matrix. (Some of the off-diagonal elements of  $R_s$  should be non zeros.) Obviously, if signal samples  $s(n)$  are i.i.d., then  $R_s = \sigma_s^2 I_L$ . In this case, the assumption is invalid, and the algorithm cannot detect the presence of the signal.

However, usually, the signal samples should be correlated due to three reasons.

1)The signal is oversampled. Let  $T_0$  be the Nyquist sampling period of signal  $s_c(t)$  and  $s_c(nT_0)$  be the sampled signal based on the Nyquist sampling rate. Based on the sampling theorem, signal  $s_c(t)$  can be expressed as

$$s_c(t) = \sum_{n=-\infty}^{\infty} s_c(nT_0)g(t-nT_0) \quad (15)$$

Where  $g(t)$  is an interpolation function. Hence, the signal samples  $s(n) = s_c(nT_s)$  are only related to  $s_c(nT_0)$ . If the sampling rate at the receiver  $f_s > 1/T_0$ , i.e.,  $T_s < T_0$ , then,  $s(n) = s_c(nT_s)$  must be correlated. An example of this is a narrow-band signal, such as the wireless microphone signal. In a 6-MHz-bandwidth TV band, a wireless microphone signal only occupies about 200 kHz. When we sample the received signal with a sampling rate of not lower than 6 MHz, the wireless microphone signal is actually oversampled and, therefore, highly correlated. The propagation channel has time dispersion; thus, the actual signal component at the receiver is given by

where  $s_0(t)$  is the original transmitted signal, and  $h(t)$  is the response of the time-dispersive channel. Since sampling period  $T_s$  is usually very small, the integration (16) can be approximated as

$$s_c(t) \approx T_s \sum_{k=-\infty}^{\infty} h(kT_s) s_o(t-kT_s) \quad (17)$$

Hence

$$s_c(nT_s) \approx T_s \sum_{k=k_0}^{K_1} h(kT_s) s_o((n-k)T_s) \quad (18)$$

Where  $[K_0T_s, K_1T_s]$  is the support of channel response  $h(t)$ , i.e.,  $h(t) = 0$  for  $t \notin [K_0T_s, K_1T_s]$ .

For the time-dispersive channel,  $K_1 > K_2$ ; thus, the signal samples  $S_c(nT_s)$  are correlated, even if the original signal samples  $S_o(nT_s)$  could be i.i.d.

1. The original signal is correlated. In this case, even if the channel is a flat fading channel and there is no oversampling, the received signal samples are correlated.

Another assumption for the algorithm is that the noise samples are i.i.d. This is usually true if no filtering is used. However, if a narrow-band filter is used at the receiver, then noise samples will sometimes be correlated. To deal with this case, we need to prewhiten the noise samples or

pretransform the covariance matrix. A method is given in the Appendix to solve this problem.

The computational complexity of the algorithm is given as follows: Computing the autocorrelations of the received signal requires about  $LN_s$  multiplications and additions. Computing  $T_1(N_s)$  and  $T_2(N_s)$  requires about  $L^2$  additions. Therefore, the total number of multiplications and additions is about  $LN_s + L^2$ .

### C. Generalized Covariance-Based Algorithms

Based on the same principle as CAV, generalized covariance-based methods can be designed to detect the signal. Let  $\varphi_1$  and  $\varphi_2$  be two nonnegative functions with multiple variables. Assume that

$$\varphi_1(a) > 0, \quad \text{for } a \neq 0 \quad \varphi_1(0) = 0$$

$$\varphi_2(b) > 0, \quad \text{for } b \neq 0 \quad \varphi_2(0) = 0$$

Then, the following method can be used for signal detection:

#### Algorithm 2: Generalized Covariance-Based Detection

Step 1) Sample the received signal, as previously described.

Step 2) Choose a smoothing factor  $L$  and a threshold  $\gamma_2$ , where  $\gamma_2$  should be chosen to meet the requirement for the probability of false alarm.

Step 3) Compute sample covariance matrix  $\hat{R}_x(N_s)$ .

Step 4) Compute

$$T_4(N_s) = \psi_2(r_m(N_s), n=1, \dots, L) \quad (19)$$

$$T_3(N_s) = T_4(N_s) + \psi_1(r_{nm}(N_s), n \neq m). \quad (20)$$

Step 5) Determine the presence of the signal based on  $T_3(N_s)$ ,  $T_4(N_s)$ , and threshold  $\gamma_2$ .

. That is, if  $T_3(N_s)/T_4(N_s) > \gamma_2$ , the signal exists; otherwise, the signal does not exist.

Obviously, the CAV algorithm is a special case of the generalized method when  $\varphi_1$  and  $\varphi_2$  are absolute summation functions. As another example, we can choose  $\psi_1(a) = a^T a$  and

$\psi_2(b) = b^T b$ . For this choice

$$T_3(N_s) = \frac{1}{L} \sum_{n=1}^L \sum_{m=1}^L |r_{nm}(N_s)|^2 \quad (21)$$

$$T_4(N_s) = \frac{1}{L} \sum_{n=1}^L |r_{nn}(N_s)|^2 \quad (22)$$

### D. Spectrum Sensing Using Multiple Antennas

Multiple-antenna systems have widely been used to increase the channel capacity or improve the transmission reliability in wireless communications. In the following, we assume that there are  $M > 1$  antennas at the receiver and exploit the received signals from these antennas for spectrum sensing. In this case, the received signal at antenna  $i$  is given by

$$H_0 : x_i(n) = n_i(n) \quad (23)$$

$$H_1 : x_i(n) = S_i(n) + n_i(n) \quad (24)$$

$$X(n) = [x_1(n) \ \dots \ x_M(n)x_1(n-1) \ \dots \ x_M(n-1) \ \dots \ x_1(n-L+1) \ \dots \ x_M(n-L+1)]^T \quad (25)$$

$$S(n) = [s_1(n) \ \dots \ s_M(n)s_1(n-1) \ \dots \ s_M(n-1) \ \dots \ s_1(n-L+1) \ \dots \ s_M(n-L+1)]^T \quad (26)$$

$$\eta(n) = [\eta_1(n) \ \dots \ \eta_M(n)\eta_1(n-1) \ \dots \ \eta_M(n-1) \ \dots \ \eta_1(n-L+1) \ \dots \ \eta_M(n-L+1)]^T \quad (27)$$

In hypothesis  $H_1$ ,  $S_i(n)$  is the signal component received by antenna  $i$ . Since all  $S_i(n)$ 's are generated from the same source signal, the  $S_i(n)$ 's are correlated for  $i$ .

It is assumed that the  $n_i(n)$ 's are i.i.d. for  $n$  and  $i$ .

Let us combine all the signals from the  $M$  antennas and

define the vectors in (25)–(27), as shown in the bottom of the page. Note that (3)–(5) are a special case ( $M = 1$ ) of the preceding equations. Defining the statistical covariance matrices in the same way as those in (6) and (7), we obtain

$$R_x = R_s + \sigma_n^2 I_{ML} \quad (28)$$

Except for the different matrix dimensions, the preceding equation is the same as (8). Hence, the CAV algorithm and generalized covariance-based method previously described can directly be used for the multiple-antenna case. Let  $S_o(n)$  be the source signal. The received signal at antenna  $i$  is

$$S_i(n) = \sum_{k=0}^{N_i} h_i(k) S_o(n-k) + n_i(n), \quad i=1,2,\dots,M \quad (29)$$

### III. COMPARISON WITH ENERGY DETECTION

Energy detection is the basic sensing method, which was first proposed in [14] and further studied in [10]–[12] and [15]. It does not need any information of the signal to be detected and is robust to unknown dispersive channels. Energy detection compares the average power of the

received signal with the noise power to make a decision. To guarantee reliable detection, the threshold must be set according to the noise power and the number of samples [10]–[12]. On the other hand, the proposed methods do not rely on the noise power to set the threshold [see (76)] while keeping the other advantages of energy detection. Accurate knowledge of the noise power is the key to the energy detection. Unfortunately, in practice, noise uncertainty is always present. Due to noise uncertainty [10]–[12], the estimated (or assumed) noise power may be different from the real noise power. Let the estimated noise power be

$\hat{\sigma}_n^2 = \alpha \sigma_n^2$  We define the noise uncertainty factor (in decibels) as

$$B = \sup \{10 \log_{10} \alpha\} \quad (34)$$

It is assumed that  $\alpha$  (in decibels) is evenly distributed in an interval  $[-B, B]$  [11]. In practice, the noise uncertainty factor of a receiving device normally ranges from 1 to 2 dB [11], [20]. Environment/interference noise uncertainty can be much higher [11]. When there is noise uncertainty, the energy detection is not effective [10]–[12], [20]. The simulation presented in the next section also shows that the proposed method is much better than the energy detection when noise uncertainty is present. Hence, here, we only compare the proposed method with ideal energy detection (without noise uncertainty).

To compare the performances of the two methods, we first need a criterion. By properly choosing the thresholds, both methods can achieve any given  $P_d$  and  $P_{fa} > 0$  if a sufficiently large number of samples are available. The key point is how many samples are needed to achieve the given  $P_d$  and  $P_{fa} > 0$ . Hence, we choose this as the criterion for comparing the two algorithms. For energy detection, the required number of samples is approximately [11]

$$N_e = \frac{2(Q^{-1}(P_{fa}) - Q^{-1}(P_d))^2}{SNR^2} \quad (35)$$

Comparing (79) and (82), if we want  $N_c < N_e$ , we need

$$\gamma_L > 1 + \frac{L-1}{\sqrt{\pi} (Q^{-1}(P_{fa}) - Q^{-1}(P_d))} \quad (36)$$

For example, if  $P_d = 0.9$  and  $P_{fa} = 0.1$ , we need

$$\gamma_L > 1 + (L-1/4.54)$$

In conclusion, if the signal samples are highly

correlated such that (83) holds, CAV detection is better than ideal energy detection; otherwise, ideal energy detection is better.

In terms of the computational complexity, the energy detection needs about  $N_s$  multiplications and additions. Hence, the computational complexity of the proposed methods is about  $L$  times that of the energy detection.

#### IV. RESULTS

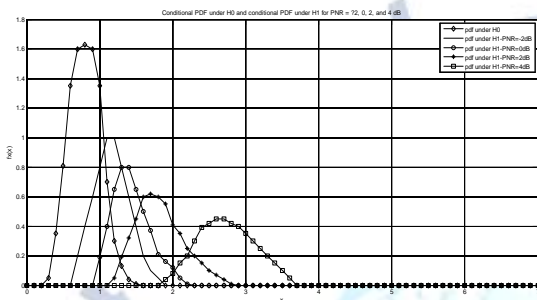


Fig. 1. Conditional PDF under  $H_0$  and conditional PDF under  $H_1$  for PNR = -2, 0, 2, and 4 dB.

From the fig 2 If the noise variance is exactly known ( $B = 0$ ), the energy detection is better than the proposed method. However, as discussed in [10]–[12], noise uncertainty is always present. Even if the noise uncertainty is only 0.5 dB, the  $P_d$  of the energy detection is much worse than that of the proposed method.

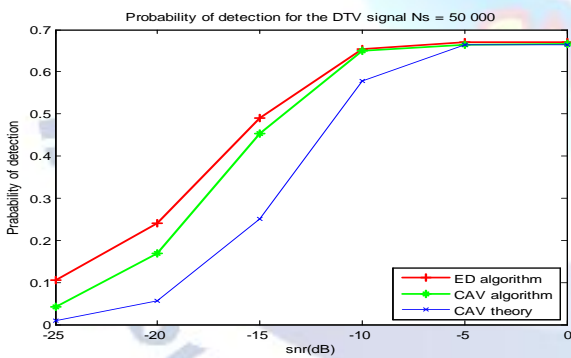


Fig 2: probability of detection for the DTV signal  $N_s = 50\ 000$ .

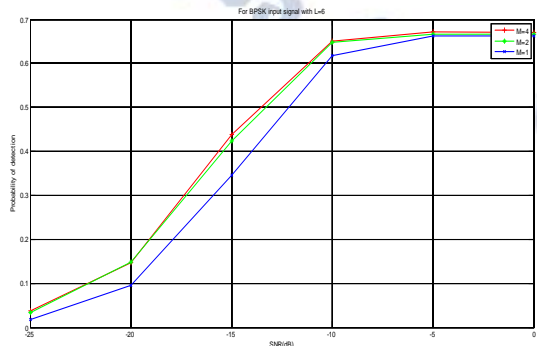


Fig. 3. Probability of detection using multiple antennas:  $P_{fa} = 0.1$ , and  $N_s = 20$ , with three source signals.

The channel taps are generated as Gaussian random numbers and different for different Monte Carlo realizations. Source signal  $s_0(n)$  is i.i.d. and binary phase-shift keying modulated. From fig 3 The received signal at antenna  $i$  is defined in (29). The smoothing factor is  $L = 6$ , and the number of samples at each antenna is  $N_s = 20$ . We fix the  $P_{fa} = 0.1$  at all cases

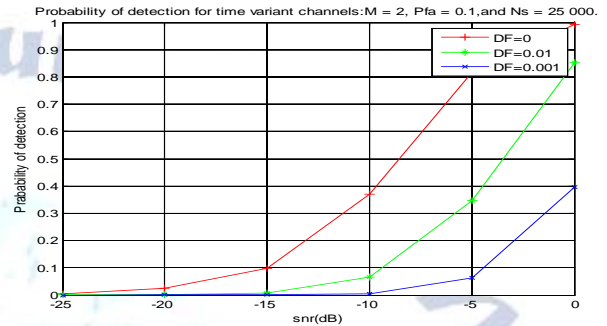


Fig 4: Probability of detection for time variant channels:  $M = 2$ ,  $P_{fa} = 0.1$ , and  $N_s = 25\ 000$ .

#### CONCLUSION

In this paper, sensing algorithms based on the sample covariance matrix of the received signal have been proposed. Statistical theories have been used to set the thresholds and obtain the probabilities of detection. The methods can be used for various signal detection applications without knowledge of the signal, channel, and noise power. Simulations based on the narrow-band signals, captured DTV signals, and multiple antenna signals have been carried out to evaluate the performance of the proposed methods. It is shown that the proposed methods are, in general, better than the energy detector when noise uncertainty is present. Furthermore, when the received signals are highly correlated, the proposed method is better than the energy detector, even if the noise power is perfectly known.

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